The relationship between parental disability and child outcomes: Evidence from veteran families

Leah K. Lakdawala 1* , Prashant Bharadwaj 2 ,

 Department of Economics, Wake Forest University, Kirby Hall Box 7505, Winston-Salem, NC 27109
 Department of Economics, University of California, San Diego, 9500 Gilman Dr., La Jolla, CA 92093-0508

 \ast Corresponding author: Leah K Lakdawala, lakdawl@wfu.edu

Abstract

We examine the relationship between parental disability and child outcomes in the American Community Survey. We focus on families with veteran parents, for whom parental disability is a direct result of service-related activities and thus is more plausibly exogenous to child outcomes than other forms of parental disability. Using the service connected disability rating (SCDR) as a measure of the severity of veteran disability, we document a gradient in child outcomes with respect to parental disability (even conditional on having a disabled parent). Children with more severely disabled parents are more likely to be late for grade, less likely to be in private school, and more likely to have disabilities themselves. These results lend meaningful insight to broader populations; we find similar associations between parental disability and child outcomes in non-veteran families. We provide evidence consistent with two broad mechanisms: first, parental disability reduces parental labor supply and thus household income (even net of transfers) and second, children — especially older children — allocate time away from work and schooling to provide care for disabled parents.

Keywords: disability, health, schooling, human capital.

1 Introduction

The question of how parental disability status affects child outcomes is of critical importance. A rich literature has shown that children's environments have large impacts on childhood development and can affect health and labor market outcomes in the long run [1,2]. Yet we know little about the consequences of parental disability for children.

Parental disability can potentially disrupt the accumulation of children's human capital in a variety of ways. First, disability is strongly correlated with poverty in both developed and developing countries [3–7]. For example, in the U.S., working age adults (ages 18-64) with a disability are more than twice as likely to be in poverty (25% compared to 9.9%).¹ Disability can simultaneously preclude parents from working² and increase household expenses if disabled parents require increased care. This reduction in financial resources can affect the health care that children in the household receive. This is important for both physical and cognitive disabilities, as early diagnosis and treatment is often critical for limiting the severity of disability [8,9]. As a result, child health is strongly related to household income [10–13].

Second, parental disability can impact the quality of the home environment. A large body of work documents the effects of children's environments on their health and educational outcomes (see, for example, [14, 15] and [16]). In the context of parental disability, one potential pathway works through the disabled parent's need for additional care. When this care is provided by non-disabled spouse, it reduces the time and resources the spouse can allocate to children [17] as well as the ability to supply labor [18]. When such care is provided by children, qualitative work has found that this has a negative effect on children's development and childhood experience [19]. Parental cognitive disabilities and injuries such as traumatic brain injury and post-traumatic stress disorder (PTSD) may be particularly harmful to children; for example, PTSD has been associated with higher levels of family violence, marital conflicts, and family distress [20].³ Hence, if parental disability places families in poverty or otherwise disrupts schooling and reduces investments in children, it could have long run consequences that may be difficult or expensive to undo.

To better understand the link between parental disability and child outcomes, this 30 paper examines the empirical relationship between parental disability among veterans 31 and child outcomes at a national level using 12 years of data from the American 32 Community Survey (ACS). Veterans form a considerable proportion of the American 33 population (around 6.8% of the adult population) and have some of the highest rates of 34 disability: according to the 2019 ACS, while around 14.6% of non-veteran adults are 35 disabled, the rate is double (30.7%) among veterans. Given the high rate of disability among veterans, it is unsurprising that children of veterans are especially vulnerable to 37 parental disability; while 8.4% of all children under the age of 18 live with at least one 38 disabled parent in the broader population, that figure jumps to 18.1% when considering children of veterans.⁴ Thus, in order to shed light on the intergenerational consequences 40 of disability in this population, it is vital to understand how veteran disability affects 41 children's health and schooling outcomes. 42

10

11

12

13

14

15

16

17

18

19

20

21

22

23

24

25

26

27

28

¹Authors' own calculations using the 2019 American Community Survey (ACS). Individuals are considered disabled if they report at least one disability in the following categories: cognitive, physical, vision, hearing, self-care or independent living difficulty. Poverty status is defined as having family income below 100% of the poverty line as reported in the ACS.

 $^{^{2}}$ A large body of economics research has found that disability negatively affects labor market outcomes (see [4] for a review of this literature).

³Additionally, [21] find that parental disability is associated with lower educational expectations (on the part of both parents and youths).

 $^{^{4}}$ Authors' calculations using the 2019 ACS, using the definition of disability above.

One of the empirical difficulties of examining the impacts of parental disability on child outcomes is that parental disability might be correlated with other attributes that could also affect child outcomes. While exogenous variation in disability status is difficult to find in the general population, we can get closer to the (statistical) ideal of random assignment of disability by examining parental disability in veteran families. Eligible veterans are assigned a "service connected disability rating" (henceforth, SCDR), which ranges from 0-100% and represents the extent of disability due to military service. Since this measure attempts to capture disability specifically due to military service (and not preexisting conditions or disability due to other sources), it is less likely to be driven by underlying unobservable factors otherwise correlated with child outcomes.⁵ For example, disabilities due to service-related injuries are less likely to reflect confounding background characteristics such as parental education, which is likely to be correlated with both parental disability and with children's schooling investments. Thus, by focusing on the sample of children living with veteran parents, we argue that conditional on having a parent that selected into military service, the degree of the parent's service-related disability is plausibly exogenous. This is similar to the source of identification used in [23], who find that wartime wounds from World War II service affected veterans' subsequent labor market outcomes and even the long-run outcomes of their adult children. Moreover, our setting allows us to compare outcomes of children with more and less severely disabled parents, conditional on having a disabled parent. This is important, as it allows us to further limit the bias that results from parents "selecting into" disability.

We find that children (aged 5-18) living with a veteran parent are significantly worse 65 off along schooling and health dimensions when their parent is severely disabled. 66 relative to children whose parent is less severely disabled and to children in families 67 where neither parent is disabled. A child whose parent has the highest disability rating 68 is 6.5% more likely to be late for grade and 48% more likely to report cognitive 69 difficulties compared to a child whose parent has no disability rating. The gradient in 70 child outcomes with respect to parental disability is very steep and outcomes for 71 children with more severely disabled parents are statistically and meaningfully different 72 than for children whose parents are less severely disabled (but are still disabled). These 73 associations point to the idea that children of more severely disabled veterans are likely 74 to enter adulthood at a disadvantage. Moreover, the level of benefits given to a veteran 75 increase as the SCDR increases; since our results do not condition on the receipt of such 76 benefits, these effects can be interpreted as the overall effect of SCDR status (including 77 any benefits that veterans receive). Therefore, despite the higher level of benefits that 78 accrue to more disabled veterans, their children appear to be worse off on the dimensions we can measure.

43

44

45

46

47

48

49

50

51

52

53

54

55

57

58

59

60

61

62

63

64

79

⁵Military disability often results from combat service, which [22] argue is conditionally exogenous for soldiers.

These findings do not appear to be driven by differences in characteristics of families with more or less severely disabled parents, nor do they seem due to differential selection into parenthood or living with children following disability. Interestingly, we find very little heterogeneity in this relationship across children's race and sex.

We also show that effects of parental disability on children are not limited to the veteran population. In order to get a sense of these relationships in a broader context, we compute correlations between parental disability status and child outcomes in non-veteran families. Within this population, we do not have data on the extent of disability and are hence constrained to only examining outcomes by parental disability status – i.e., whether or not a parent in the household is disabled. Thus we compare correlations between an indicator for parental disability and child outcomes across veteran and non-veteran families. We find that the negative associations between parental disability status and child outcomes are similar for veteran and non-veteran families, though the relationship is stronger in the non-veteran population, where disability is less likely to be exogenous with respect to children's outcomes.

We also provide evidence that the adverse consequences of parental disability operate 96 in part through two broad channels. First, we show that household income per capita 97 declines sharply with parental disability. This is driven by large reductions in a parent's 98 labor supply and earnings as that parent's disability is more severe. Though Veterans 99 Administration (VA) and Supplemental Security Income (SSI) transfers increase with 100 parental disability, they are not enough to offset the total decline in household earned 101 income per capita; thus, on net, children of more severely disabled parents live in poorer 102 households. Second, we illustrate that teens - who are more likely to be capable of 103 providing care for disabled parents compared to younger siblings – are less likely to 104 work when their parents are more severely disabled. Moreover, we find that working 105 teens take jobs that involve shorter commutes when their parents are severely disabled. 106 These findings are both consistent with parental disability requiring care that is often 107 provided by older children in the household. Indeed, the negative relationship between 108 work and schooling outcomes is concentrated among high school-aged children (ages 109 14-18) and in families where the parental disability explicitly requires care. 110

To our knowledge, ours is the first paper to quantify the relationship between the degree of parental disability and childhood wellbeing.⁶ Recent work has illustrated that parental disability can have long-reaching effects on the adult socioeconomic status and mortality of the next generation [23], and our results provide evidence for pathways through which these effects arise; children of severely disabled parents are at a disadvantage at early ages along the dimensions of schooling and health. In this way,

81

82

83

84

86

87

89

91

92

93

⁶Several studies compare that children of disabled parents to those of non-disabled parents and find that children of disabled parents tend to fare worse on a range of outcomes. However, these studies focus only on comparing children of disabled versus non-disabled parents (rather than the gradient of child outcomes with respect to the severity of parental disability) and the degree to which they can address the endogeneity of parental disability varies across studies (see [24], [21] and [25] for the US and [26] for Vietnam).

our results are useful for understanding the intergenerational transmission of health shocks more generally.⁷

Our paper also contributes to the literature on the determinants of childhood disability (see [28] for a summary of some recent literature). Specifically, our results illustrate that parental disability is one important factor that influences child disability, over and above potential genetic transmission of disability. [29] estimate that a family with a disabled child faces disability-associated costs of \$30,500 per year and 7.1 million children (14%) in public schools receive special education services costing about \$50 billion annually.⁸ Given the importance of childhood disability to families and to governments, understanding the link between childhood and parental disability is critical.

Finally, the results in this paper are relevant to the broader research agenda that seeks to understand the effects of parental and family health shocks. A vast body of previous work has established that parental illness and death have critical effects on the wellbeing of the household and household members (see, for example, [30], [31], [32], and [33]).⁹ Our findings add to this literature by showing that parental disability resulting from military service acts as a significant shock to veteran households and accordingly disrupts the accumulation of children's human capital.

2 Data

The results we present in this paper use data from the American Community Survey 136 (ACS) for the years 2008-2019, the years in which the question of service connected 137 disability rating (SCDR) is asked of veterans.¹⁰ The ACS is a 1-in-100 national random 138 sample of the population. One adult from each household responds to survey questions 139 on behalf of all household members, including children. Our main sample is formed of all 140 children of the household head between the ages of 5 and 18 (inclusive) who reside with 141 at least one veteran parent. This leaves us with over 481,000 children across 12 survey 142 years. For the analysis that uses non-veteran families, our sample includes all children of 143 the household head between the ages of 5 and 18, resulting in over 5.1 million children. 144

Our main covariate of interest is parental SCDR. SCDR "connotes many factors but basically it means that the facts, shown by evidence, establish that a particular injury or disease resulting in disability was incurred coincident with service in the Armed Forces, or if preexisting such service, was aggravated therein" (38 CFR 3.303).¹¹ The

117

118

119

120

121

122

123

124

125

126

127

128

129

130

131

132

133

134

⁷There is a vast literature documenting the intergenerational transmission of socioeconomic status and health; for example, see [27].

⁸American Institutes for Research (2004) https://www.air.org/sites/default/files/ SEEP1-What-Are-We-Spending-On.pdf and National Center for Education Statistics (2020) https://nces.ed.gov/programs/coe/indicator_cgg.asp.

⁹More recent work also shows that siblings' disability can impact children's schooling decisions [34]. ¹⁰We accessed the data through IPUMS USA [35]. The underlying data source is the U.S. Census Bureau. The analysis in this paper complies with the IPUMS USA terms and conditions.

¹¹Conditions that determine eligibility typically exclude, "the result of the veteran's own willful

SCDR is meant to represent a composite measure of both the severity and the connectedness of the disabilities to service. This score is typically calculated when a veteran applies for disability compensation after having undergone a medical exam at a VA hospital. The SCDR reflects both physical disabilities (such as amputations or sensory impairment) and non-physical disabilities (such as post-traumatic stress disorder, PTSD). The SCDR and household demographics determine the level of benefits for which a veteran is eligible; these benefits generally increase linearly with SCDR with the exception of benefits tied to SCDRs of 100%, which are much more generous (see S1 Fig).¹² Though the score is reported to veterans and relevant administrators in increments of 10 percentage points, in the ACS we observe the SCDR only in bins of 20 percentage points and it is top-coded at 70 percent.

We also observe some dimensions of general self-reported disability in the ACS for 160 all individuals, including non-veterans. These are indicators for whether an individual 161 has any (i) cognitive difficulties – difficulty learning, remembering, concentrating, or 162 making decisions because of a physical, mental, or emotional condition; (ii) physical 163 difficulties – limitations on "basic" physical activities, such as walking, climbing stairs, 164 reaching, lifting, or carrying; (iii) "long-lasting" condition of blindness, deafness, or a 165 severe vision or hearing impairment; and (iv) self-care and independent living difficulties 166 - inabilities to care for oneself (not including temporary health conditions such broken 167 bones or pregnancy) either within (self-care) or outside (independent living) the home. 168 The ACS does not contain information on the severity or number of disabilities.¹³ 169

In our sample, self-reported disability and SCDR are highly correlated among 170 veteran parents (Fig 1(a)-(d)). Self-reported disability across all categories is increasing 171 in SCDR, with a discrete jump up at SCDRs of 70 percent or higher (potentially due to 172 the top-coding of SCDRs in the ACS). The likelihood of reporting any disability and 173 the number of reported disability categories is also increasing in SCDR (Fig 1(e)) and 174 Fig 1(f). On average, those with an SCDR of 70% or higher have disabilities that span 175 more than one category. However, self-reported disability does not perfectly correspond 176 with SCDRs. For example, about 10% of parents without any SCDR report disability 177 (Figure Fig 1(e)); this is because SCDRs apply only to disabilities sustained or worsened 178 due to military service and thus excludes non-service-related disabilities. Additionally, 179 not all individuals with SCDRs self-report disabilities. This could be because the 180

149

150

151

152

153

154

155

156

157

158

misconduct or, for claims filed after October 31, 1990, the result of his or her abuse of alcohol or drugs" (38 U.S.C. 105).

¹²Veterans with certain severe disabilities or disabilities with "special circumstances such as the need of aid and attendance by another person or by specific disability, such as loss of use of one hand or leg" may be eligible for additional special monthly compensation (SMC) "paid based on the need of aid and attendance by another person" (as reported by the VA, http://www.benefits.va.gov/). Additionally, surviving dependents of veterans who died due to service-related disabilities are eligible to receive Dependency and Indemnity Compensation (DIC). In this analysis we restrict our attention to children in households with living but disabled veterans, so the large majority will not be eligible for DIC.

¹³We do not observe the total number of disabilities in the ACS, only the number of disability categories that apply to each individual. It is possible for a person to have multiple disabilities that fall into the same category.

categories of self-reported disability in the ACS do not cover all disability types. 181 Nonetheless, given the strong correlations between self-reported disability measures (including self-reporting disabilities in multiple categories) and SCDR presented in these 183 figures, we believe that variation in parental SCDR captures a combination of the 184 likelihood and severity of parental disability. 185

Fig 1. SCDR and Self-Reported Disabilities among Veterans.

Tables 1 and 2 display summary statistics for the samples of children of veteran and 186 non-veteran families. We present statistics for both samples because, as discussed in 187 Section 3, we examine correlates of parental disability for both samples of children. 188 Column (1) shows that about 19% of children have a parent with an SCDR, conditional 189 on living with a veteran parent.¹⁴ About 7% of children live with a veteran with an 190 SCDR of 10-20% (which we refer to as having a less severely disabled parent from this 191 point onwards), and 7% of children live with a parent with the highest SCDR rating 192 (70% and above), which we refer to as having a "severely disabled" parent from this 193 point onwards. Self-reported parental disability is 16% in the veteran family sample 194 (column (1)), which is significantly higher than in the non-veteran sample (9%, column 195 (2)).196

Children of veterans appear to be very similar to children of non-veterans along most 197 dimensions (e.g. in terms of sex, age, birth order, race, household size, number of 198 siblings, and father's age). It is worth noting that because of the large sample size, even 199 small differences - which are not economically meaningful - are statistically significant, 200 so we interpret the p-values reported in column 3 with caution. However, there are a 201 few key differences. Children of veterans have slightly older parents and are less likely 202 to be missing information on father's education than children of non-veterans. This is 203 because children of veterans are much more likely to be living with their fathers (96.2%)204 versus 81.1%). They also have higher household income per capita (by about \$1400).¹⁵ 205

The main child outcomes we consider are current schooling, late-for-grade, work, and 206 child disability status across a number of indicators. "Currently in School" is an 207 indicator that is equal to one if a child has attended a school within the last three 208 months. Additionally, for children who are currently attending school, we observe 209 whether the school they attend is public or private.¹⁶ As expected, most children are in 210 school (96.5%); about 12.2% are in private school (Table 2). We use private school 211 attendance as a very broad measure of schooling investment/outcomes, for two reasons. 212 First, children who attend private schools (whether due to selection or quality) perform 213 better than students who attend public schools, along standard dimensions of 214

 $^{^{14}\}mathrm{In}$ the case where we observe both parents with an SCDR, we use the higher SCDR.

¹⁵Household income per capita is winsorized at the 99.5th percentile within each state and survey year and is expressed in 1999 dollars using the CPI-U multiplier published by the Bureau of Labor Statistics. ¹⁶In the ACS, home-schooling is included under the classification of private schooling as is not separately identifiable.

1		•	
	Children	Children	p-value for
	with at least	with	$H_0:(1) =$
	one veteran	non-veteran	(2)
	parent	parents	
	(1)	(2)	(3)
Household Size	4.49	4.50	0.001
	[1.38]	[1.48]	
Number of Siblings in HH	1.45	1.57	0.000
	[1.2]	[1.24]	
Number Grandparents in HH	0.04	0.04	0.000
	[0.21]	[0.23]	
Mother's Age	41.65	40.40	0.000
-	[7.39]	[7.21]	
Mother's Education			
High School or Less	0.30	0.38	
1 Year of College	0.17	0.14	<i>p</i> - value for the joint test that distribution is
2 Years of College	0.13	0.10	the same across groups
4 or More Years of	0.34	0.34	= 0.000
College			
Missing	0.05	0.04	
Father's Age	44.99	42.99	0.000
C C	[8.73]	[7.75]	
Father's Education			
High School or Less	0.34	0.35	
1 Year of College	0.20	0.10	<i>p</i> - value for the joint test that distribution is
2 Years of College	0.12	0.06	the same across groups
4 or More Years of	0.30	0.29	= 0.000
College			
Missing	0.04	0.19	
Any Parental Disability	0.16	0.09	0.000
Parental SCDR			
No Disability	0.78		
Rating			
10 to 20 percent	0.07		
30 to 40 percent	0.04		
50 to 60 percent	0.03		
70 percent or more	0.07		
Household Income Per Capita	17742.55	16347.67	0.000
1	[15062.6]	[17814.79]	
Number of observations	481,725	5,126,450	

Table 1. Descriptive Statistics for Children Ages 5-17: Child Characteristics

Data from the American Community Survey (2008-2015). Standard deviations in square brackets below means. Household income per capita trimmed at the bottom and top 1% within each survey year and is expressed in 1999 dollars using the CPI-U multiplier published by the Bureau of Labor Statistics. Family income as a percentage of the poverty line is as reported in the ACS, which uses the poverty line established the Social Security Administration in 1964 and subsequently revised in 1980 (adjusted for inflation) as well as detailed income and family structure information. Column 3 reports the p-value for the test that the means across veteran and non-veteran samples are the same. However, due to large sample sizes, the p-values are almost always 0, even when the difference in means is not economically meaningful. Thus, we interpret these p-values with caution.

performance such as test scores [36–40] as well as high school graduation and college 215 attendance [41–43]. Second, private schools are typically more expensive than public 216 schools. According to the most recent report from the National Center for Education 217 Statistics, private school tuition is on average \$12,420 per year [44]. Moreover, [45] find 218 that the propensity to attend private school increases with both income and ability, 219 suggesting that private school attendance could reflect greater investment in schooling, 220 higher ability (which could be affected by a parents' disability status), or both. We 221 classify 4.6% of the sample as "Late for Grade", which we define as being at least 2 222 years below the modal grade-for-age in the ACS. This measure is meant to capture 223 slower-than-normal progression through school grades and possible grade retention.¹⁷ 224

There are several labor force outcomes we consider, all of which are only reported for those aged 16 and older.¹⁸ The first is an indicator for whether an individual has been employed in the previous year. Around 27% of the sample ages 16-18 report being employed in the past year at an average of 460 hours per year (about 8.8 hours per week). Earnings are on average close to \$7 per hour.¹⁹

Child disability status is reported in the survey along the same dimensions as for parents. Cognitive difficulties are the most common in our sample (Table 2), affecting 4.4% of children. Physical as well as "long-lasting" sensory (i.e., vision and hearing) difficulties are much less common (0.6%-1.2% of children). About 1% of the sample suffers from self-care difficulty, while about 2.2% of the sample (aged 15 or older) suffers from independent living difficulties.

225

226

227

228

229

230

231

232

233

234

 $^{^{17}}$ We calculate this variable for children 8 and older; compulsory school starting ages are state-specific and vary from age 5 to age 8 (as reported in 2008 by the U.S. Department of Education, Institute for Education Sciences, National Center for Education Statistics (http://nces.ed.gov/)).

¹⁸We do not study children over the age of 18, as the survey only contains information for coresident parents. Thus as age increases, the sample of individuals for which we observe parental disability is likely to become less and less representative of the population.

¹⁹Hourly earnings are calculated using reported hours worked in a "usual" week and weeks worked in the past year. However, weeks worked in the previous year are only reported in intervals so the midpoint of each interval is used as is standard [46]. Hourly earnings are winsorized at the 99.5th percentile (across all workers) within each state and survey year and are expressed in 1999 dollars using the CPI-U multiplier published by the Bureau of Labor Statistics.

	Children	Children	p-value for
	with at least	with	$H_0:(1) =$
	one veteran	non-veteran	(2)
	parent	parents	
	(1)	(2)	(3)
Female	0.49	0.49	0.135
Age	11.97	11.52	0.000
	[3.93]	[3.95]	
Birth Order	1.70	1.73	0.000
	[0.9]	[0.93]	
White	0.73	0.62	0.000
Black	0.12	0.11	0.000
Hispanic	0.12	0.20	0.000
Schooling and Labor Force Outcomes			
In School (Previous 3 Months)	0.965	0.965	0.014
Attending Private School	0.122	0.126	0.000
Late for Grade	0.046	0.047	0.000
Employed (Previous Year)	0.269	0.249	0.000
Hours Worked (Previous Year)	460.2	452.9	0.002
	[463.15]	[484.07]	
Hourly Earnings (Previous Year)	6.85	7.18	0.000
	[8.57]	[9.23]	
Disabilities			
Cognitive Difficulties	0.044	0.038	0.000
Physical Difficulties	0.006	0.006	0.066
Sensory Difficulties	0.012	0.013	0.002
Self-care Difficulty	0.010	0.009	0.013
Independent Living Difficulty	0.025	0.022	0.000
Any Difficulty	0.057	0.051	0.000
Number of observations	481,725	$5,\!126,\!450$	

Table 2. Descriptive Statistics for Children Ages 5-17: Household Characteristics

Data from the American Community Survey (2008-2015). Standard deviations in square brackets below means. Employment (and thus normalized earnings) information only asked of individuals aged 16 or older. Independent living difficulty is only asked of individuals aged 15 or older. Hourly earnings are trimmed at the bottom and top 1% within each survey year and is expressed in 1999 dollars using the CPI-U multiplier published by the Bureau of Labor Statistics. Column 3 reports the p-value for the test that the means across veteran and non-veteran samples are the same. However, due to large sample sizes, the p-values are almost always 0, even when the difference in means is not economically meaningful. Thus, we interpret these p-values with caution.

3 Empirical Approach and Results

3.1 The relationship between the degree of parental disability 237 and child outcomes in veteran families 238

Our baseline estimates are generated by running the following regression on the sample ²³⁹ of children age 5-17 with at least one veteran parent: ²⁴⁰

$$Y_{iht} = \beta_1 \cdot \mathbf{1}(SCDR_{ht} = 10 \text{ or } 20 \text{ percent})$$
(1)
+ $\beta_2 \cdot \mathbf{1}(SCDR_{ht} = 30 \text{ or } 40 \text{ percent}) + \beta_3 \cdot \mathbf{1}(SCDR_{ht} = 50 \text{ or } 60 \text{ percent})$
+ $\beta_4 \cdot \mathbf{1}(SCDR_{ht} = 70 \text{ percent or higher}) + \gamma X_{iht} + \delta_t + \mu_s + \varepsilon_{iht}$

where Y_{iht} is an outcome of interest such as schooling or disability status for child *i* 241 in household h in survey year t; $SCDR_{ht}$ is the parental veteran disability score²⁰ in 242 increments of 20 percentage points and top-coded at 70 percent; X_{iht} are child- and 243 household-level characteristics such as age, race, and household size²¹; and δ_t and μ_s 244 represent survey year and state fixed effects that capture aggregate differences in Y_{iht} 245 by year and across states. β_1 , β_2 , β_3 , and β_4 capture the difference in Y_{iht} relative to 246 children of veterans without an SCDR, i.e. those who are less likely to be and/or who 247 are less severely disabled. In practice, many of the outcome variables we consider are 248 binary, so in many cases (1) represents a linear probability model. As we sometimes 249 observe multiple children of the same set of parents, we cluster standard errors at the 250 family level. 251

Though SCDRs are not randomly assigned, we believe that the sample and type of 252 disability under study help move us closer to causal estimates than existing work for 253 several reasons. First, unlike other types of injuries, injuries sustained during military 254 service captured through the SCDR are likely to be unanticipated and unrelated to 255 most preexisting health measures (recall that the score only reflects injuries sustained 256 during service and worsening of preexisting conditions due to service). Therefore, unlike 257 other measures of disability, the SCDR is likely to capture plausibly exogenous variation 258 in parental disability. Second, the SCDR allows us to examine the gradient of outcomes 259 with respect to a measure of the *severity* of parental disability. Specifically, by 260 comparing $\hat{\beta}_2$, $\hat{\beta}_3$, and $\hat{\beta}_4$ to $\hat{\beta}_1$, we can better understand how the degree of parental 261 disability matters for child outcomes, conditional on having a parent with at least some 262 degree of disability. This helps account for potential parental selection into disability. It 263

 $^{^{20}\}mathrm{As}$ noted in Section 2, in cases where both parents report an SCDR we use the higher of the two scores.

²¹Importantly, X_{iht} contains the demographic information used to determine the VA benefit eligibility. Complete list of controls: age FE, gender, dummy variables for single race categories (white, black, Hispanic), household size FE, FE for birth order, FE for number of siblings, FE for number of grandparents in household, mother's and father's age, education, and marital status FE as well as indicators for whether mothers and father served in 2001 and later (including indicators for missing parental information), FE for metro status, state FE, survey year FE.

	In School Late for Grade		In Private School
	(All Ages 5-18)	(Ages 7-18 & In School)	(Ages 5-18 & In School)
	(1)	(2)	(3)
Parental SCDR			
10 to 20 Percent	0.001	-0.000	-0.002
	(0.001)	(0.001)	(0.002)
30 to 40 Percent	-0.001	-0.001	-0.012***
	(0.001)	(0.002)	(0.003)
50 to 60 Percent	0.000	-0.002	-0.018***
	(0.002)	(0.002)	(0.003)
70 Percent or Higher	0.001	0.003**	-0.016***
	(0.001)	(0.002)	(0.002)
Observations	481,725	415,078	465,053
Mean of dependent variable	0.965	0.0461	0.125
p-value for test that			
SCDR 10-20=SCDR 70+	0.630	0.0970	0.000

Table 3. Degree of Parental Disability and Schooling Outcomes for Children of Veterans

*** p<0.01, ** p<0.05, * p<0.1 Standard errors clustered at the household level. Omitted group: Children in families where neither parent has a disability rating (SCDR=0). Sample for column (1): all children ages 5-18 living with a veteran parent; sample is restricted to children age 7-18 and currently in school in column 2 and age 5-18 and currently in school in column 3. Controls: age FE, gender, dummy variables for single race categories (white, black, Hispanic), household size FE, FE for birth order, FE for number of siblings, FE for number of grandparents in household, mother's and father's age, education, and marital status FE as well as indicators for whether mothers and father served in 2001 and later (including indicators for missing parental information), FE for metro status, state FE, survey year FE. Mean is reported for children in families where neither parent has an SCDR.

is also worth highlighting that by restricting our sample to children in veteran families (i.e., with at least one veteran parent), we circumvent the issue of selection into military service. 266

Results. Table 3 presents the estimates obtained from Eq 1 when we consider 267 schooling outcomes. There is no systematic relationship between whether a child is 268 currently in school (as of the previous 3 months) and parental SCDR; the point 269 estimates are all very close to zero and precisely estimated. This is perhaps 270 unsurprising, as most the overwhelming majority of children are attending school 271 (96.5%). Conditional on being enrolled in school, children of highly disabled veterans 272 (SCDR \geq 70 percent) are 0.3 percentage points more likely to be late for grade (about 273 (6.5%) relative to children of non-disabled parents (significant at the 5% level). They are 274 also more likely to be late for grade that children with of less severely disabled parents 275 (SCDR = 10-20 percent) and this difference is statistically significant at the 10% level. 276 Children in school are also significantly less likely to be in private school when their 277 parents are more severely disabled and these differences are large relative to the 278 proportion of children in private school for non-disabled parents (column 3). For 279 example, children whose parents have the highest SCDR are 1.6 percentage points less 280 likely to attend private school (12.8%) than those with non-disabled parents, and this 281 difference is significant at the 1% level; the same is true when comparing children with 282 severely disabled parents to children with less severely disabled parents, indicating that 283 even within children of disabled veterans, the severity of parental disability matters. To 284 the extent private school captures current and past investment in education, it appears 285 that children with more disabled parents receive lower schooling investment. 286

In Table 4, we show that children of more disabled veterans are also more likely to 287

					Independent	
	Cognitive	Physical	Sensory	Self-Care	Living	Any
	Difficulty	Difficulty	Difficulty	Difficulty	Difficulty	Difficulty
	(1)	(2)	(3)	(4)	(5)	(6)
Parental SCDR						
10 to 20 Percent	0.004^{***}	-0.000	0.002^{**}	0.001	0.001	0.005^{***}
	(0.001)	(0.000)	(0.001)	(0.001)	(0.002)	(0.001)
30 to 40 Percent	0.007^{***}	0.000	0.002^{*}	0.001	-0.001	0.006^{***}
	(0.002)	(0.001)	(0.001)	(0.001)	(0.002)	(0.002)
50 to 60 Percent	0.008***	0.000	0.003**	0.001	0.002	0.010***
	(0.002)	(0.001)	(0.001)	(0.001)	(0.002)	(0.002)
70 Percent or Higher	0.020***	0.002^{***}	0.006^{***}	0.003***	0.009^{***}	0.025^{***}
	(0.002)	(0.001)	(0.001)	(0.001)	(0.002)	(0.002)
Mean of dependent variable	0.0420	0.00597	0.0114	0.00959	0.0243	0.0543
p-value for test that						
SCDR 10-20=SCDR 70+	0.000	0.000	0.000	0.002	0.000	0.000
Observations	481,725	481,725	481,725	481,725	155,927	481,725

Table 4. Degree of Parental Disability and Disability in Children of Veterans

*** p<0.01, ** p<0.05, * p<0.1 Standard errors clustered at the household level. Omitted group: Children in families where neither parent has a disability rating (SCDR=0). Sample for column 1-4 and 6: all children ages 5-18 living with a veteran parent; column 5 includes only children aged 15 or older. Controls: age FE, gender, dummy variables for single race categories (white, black, Hispanic), household size FE, FE for birth order, FE for number of siblings, FE for number of grandparents in household, mother's and father's age, education, and marital status FE as well as indicators for whether mothers and father served in 2001 and later (including indicators for missing parental information), FE for metro status, state FE, survey year FE. Mean is reported for children in families where neither parent has an SCDR.

> suffer from disabilities themselves, across a variety of disability types. Children of more 288 severely disabled veterans are significantly more likely to have cognitive difficulties, 289 defined as "serious difficulty concentrating, remembering, or making decisions" due to a 290 "physical, mental, or emotional condition" (column 1). This difference grows with the 291 severity of parental disability, as measured by parental SCDR; children of highly 292 disabled veterans (SCDR \geq 70 percent) are 2 percentage points more likely to have a 293 cognitive disability than those without disabled parents, representing a 48% increase in 294 the prevalence of cognitive difficulties among children. This finding seems especially 295 relevant, as one important way in which parental disability could affect children is 296 through by disrupting the home environment, which has a potentially meaningful 297 impact on children's mental and emotional development. There is some evidence that 298 parental disability is related to other child disabilities as well; children living with 299 parents with the highest SCDR category are more likely to suffer from disabilities in all 300 categories (columns 1-3). Given this, it is perhaps not surprising that children are less 301 able to care for themselves (column 4) or live independently (column 5; only for those 302 ages 15 and older) when their parents have SCDRs of 70 percent or more. These final 303 two categories of child disability potentially represent the most severe forms of disability 304 reported in the ACS, as they correspond to physical or mental health conditions lasting 305 6 months or longer that make it difficult for individuals to "take care of their own 306 personal needs, such as bathing, dressing, or getting around inside the home" [35]. In 307 column 6, we use a composite measure of child disability - whether the child reports a 308 disability in any category - and find a steep gradient in parental disability. Therefore 309 along the dimension of child disability, it again appears that children of highly disabled 310 veterans are at a disadvantage. 311

Robustness Checks & Other Results. To lend credibility to our assumption that 312 parental SCDR is exogenous in the population of children living with a veteran parent, 313 we show in S1 Table Table 1 that important household characteristics are largely 314 balanced across SCDR. There are no apparent systematic patterns in household size, 315 number of grandparents in the household, metro status, mother's education, child 316 gender, or home language with respect to parental SCDR (columns 1-4 and 6-7). 317 Though some coefficients are statistically significant, the point estimates tend to be 318 small and do not yield a clear pattern – notably, they do not indicate that children of 319 severely disabled parents are different along observable dimensions – and they are not 320 generally jointly significant at conventional levels. In column 5, we see that children of 321 more severely disabled parents have slightly younger mothers, but the difference is very 322 small relative to the average age of mothers; put another way, it seems unlikely that 323 having a mother who is 41 years and 8 months old versus 41 years and 9.7 months old 324 could drive the differences that we find. 325

Another potential concern is that disability could affect selection into parenthood. 326 i.e., parents that choose to have children following a disability could be different than 327 parents who choose to have children in the absence of a disability. To address this 328 concern, we use the (very limited) information we have on the timing of disability and 329 child age. In S2 Table, we restrict the sample to children who were born before their 330 parents' last tour of duty. Specifically, we include only children who were born before 331 2001, but whose parents served in the military in 2001 or later. Because we use 332 information on each individual parent's theater of war^{22} , we present the results 333 separately for children with a veteran father (columns 1 and 3) and with a veteran 334 mother (columns 2 and 4). We note that these restricted samples are considerably 335 smaller than our main samples (less than 9% of the main sample for fathers and less 336 than 2% for mothers), so this analysis is not well powered to detect the effect sizes in 337 Tables 1 and 2. That said, even for these smaller, restricted samples – where selection 338 into parenthood following disability is very unlikely – we find that children of more 339 severely disabled parents are significantly and substantively less likely to be in private 340 school, with the exception of column 2, where the point estimate is the same as in our 341 main samples but where we lack precision for statistical significance due to the small 342 sample size. For the outcome of child disability (columns 3 and 4), the effect sizes in the 343 restricted samples are statistically significant only for the highest category of parental 344 disability (70 percent or higher). Altogether, we take the results in S2 Table as 345 suggestive evidence that selection into parenthood does not fully explain our main 346 findings. 347

In order to be able to link children directly to parental disability status, in our main 348 sample we focus only on children of the household head who are currently residing with 349

 $^{^{22}\}mathrm{We}$ do not observe specific dates of service for each parent. Instead, we observe only the general "theatre of war." All veterans that have served since 2001 are grouped in a single "Global War on Terrorism / post-2001" theater.

a veteran parent. To show that our results are robust to using a less restrictive sample, 350 we run our estimation on the sample of all children currently residing with an adult 351 veteran in S3 Table. In this sample, we examine the effect of any coresiding adult's 352 SCDR on the outcomes of all children in the household (regardless of familial 353 relationship between the veteran and child(ren)). We find very similar patterns with 354 respect to private school status (column 1) and child disability status (column 2). 355 Children living with more severely disabled adults are significantly and considerably 356 more disadvantaged on these margins. 357

Parental disability could potentially lead to sample selection if disabled veteran parents are more or less likely to live apart from their children than non-disabled veteran parents because we are only able to match children to their parents' disability status if they live with their parents. To show that this type of sample selection is not driving our results, we examine the relationship between SCDR and living with children (ages 0-18) in the sample of veterans that are likely to be parents. Specifically, in S4 Table we show that, among veterans age 19-50, there is no systematic or statistically significant relationship between own SCDR and the number of coresiding children. Thus, we find no evidence of this type of sample selection.

Finally, we examine the effects of parental disability across child race and gender. Generally, S5 Table illustrates that there are few statistically significant differences in the effects by race (columns 1-2 and 5-6) or sex (columns 3-4 and 7-8). It does seem that the effects of severe parental disability (SCDR of 70 percent or more) are slightly larger for white children (columns 1 and 5) and for boys (columns 4 and 8), though the effects of other SCDR categories seem similar across age and sex.

The biggest differences arise when looking at the relationship between child 373 outcomes and mother's versus father's SCDR (S6 Table). For private school status, the 374 magnitudes are much higher for father's SCDR than for mother's SCDR (columns 1-2). 375 Conversely, the effects of mother's SCDR are much stronger for children's disability 376 status (coliumns 3-4). This pattern is consistent with the interpretation that father's 377 disability affects child schooling through an income channel, which affects the budgetary 378 aspects of schooling decisions (e.g. private versus public education), while maternal 379 disability affects child outcomes through other channels, such as mother's time 380 allocation or influence on the home environment.²³ We discuss mechanisms more 381 formally in section 4. 382

3.2 The relationship between parental disability status and child outcomes in the wider population

The findings in Section 3.1 indicate that among children in veteran families, more severe parental disability is associated with poorer outcomes in terms of schooling and the 386

358

359

360

361

362

363

364

365

366

367

368

369

370

371

372

383

 $^{^{23}}$ For example, [17] find that maternal disability lowers parents' school involvement and is associated with a less enriching home environment.

incidence of child disabilities. However, it is important to understand whether this relationship is specific to veteran families, who are observably different from the wider population (see Tables 1 and 2). While we do not observe markers for the degree of parental disability in the non-veteran sample, we do observe the existence of (self-reported) parental disability along the dimensions described in Section 2. In this section, we examine the relationship between child outcomes and parental disability status as captured by the an indicator for any parental disability (i.e. any category of disability for either parent), both in the sample of veteran families and the sample of non-veteran families. To assess whether the relationship between parental disability and child outcomes varies across these sub-populations, we also report the p-value for the test that this relationship is the same across the two samples.

It is important to keep in mind that when we look at the gradient of child outcomes with respect to SCDR in the sample of veteran families, we argue that the degree of parental disability in this subsample is likely to be exogenously determined and unlikely to capture other determinants of child outcomes (e.g. disadvantages that pre-date disability). In the wider sample and using the indicator for self-reported parental disability, this is less likely to hold. For example, the intergenerational correlation of disability in the wider population could reflect causal pathways such as time allocation and household income (as will be discussed in Section 4) or simply compositional differences across children with and without disabled parents. Thus we underscore that these estimates reflect associations rather than causal effects of parental disability. Nonetheless, as they are the first estimates of the relationship between schooling outcomes, child disability outcomes, and parental disability using a large, nationally representative dataset (that we are aware of), we still see them as an important step forward in our understanding of the intergenerational effects of parental disability.

The association between schooling outcomes and parental disability in the full sample of children ages 5-18 are displayed in Table 5. On average, children who have a disabled parent are significantly less likely to be in school, more likely to be late for grade, and less likely to be in private school (conditional on being in school). The magnitudes of these associations are meaningful. For example, children of disabled parents are nearly 36-40% more likely to be late for grade (column 2) and 11-12% less likely to attend private school (column 3). This holds true for both veteran and non-veteran families; the point estimates are similar across the two subpopulations, though they are statistically significantly different at the 5% level for late for grade status (but not for private school attendance).

Child and adult disability are also strongly correlated in the wider sample (Table 6). 422 Children are much more likely to have disabilities of all types when they have at least 423 one parent with a disability. The coefficients are large and meaningful. Having a 424 disabled parent increases the chances of a child disability by 1.1 to 3.7 times (i.e., 110% 425 to 370% over the incidence of child disability in the population of children without 426 disabled parents). Interestingly, the correlation is notably stronger in non-veteran 427 families than in veteran families (differences are statistically significant across all 428

387

388

389

390

391

392

393

394

395

396

397

398

399

400

401

402

403

404

405

406

407

408

409

410

411

412

413

414

415

416

417

418

419

420

PANEL A: Children with a veteran	parent		
	In School	Late for Grade	In Private School
	(Previous 3 Months)	(Conditional)	(Conditional)
	(1)	(2)	(3)
Parent declares a disability	-0.006***	0.015***	-0.015***
	(0.001)	(0.001)	(0.002)
Observations	481,725	415,078	465,053
Mean of dependent variable	0.967	0.0421	0.127

Table 5. Parental Disability Status and Schooling Outcomes for Children

PANEL B: Children with non-veteran parents

	In School	Late for Grade	In Private School
	(Previous 3 Months)	(Conditional)	(Conditional)
	(1)	(2)	(3)
Parent declares a disability	-0.006***	0.018***	-0.014***
	(0.000)	(0.000)	(0.001)
Observations	$5,\!126,\!450$	4,303,343	4,945,546
Mean of dependent variable	0.966	0.0447	0.129
p-value for H_0 : Panel A = Panel B	0.898	0.0208	0.324

**** p<0.01, ** p<0.05, * p<0.1 Standard errors clustered at the household level. Sample for column (1): all children ages 5-18; sample is restricted to children age 7-18 and currently in school in column 2 and age 5-18 and currently in school in column 3. Controls: age FE, gender, dummy variables for general race categories (white, black, other - omitted; Hispanic - nonexclusive), household size FE, FE for birth order, FE for number of siblings, FE for number of grandparents in household, mother's and father's age and education FE (including indicators for missing parental information), FE for metro status, state FE, survey year FE.

PANEL A: Children with a veteran parent

					Independent	
	Cognitive	Physical	Sensory	Self-Care	Living	Any
	Difficulty	Difficulty	Difficulty	Difficulty	Difficulty	Difficulty
	(1)	(2)	(3)	(4)	(5)	(6)
Parent declares a disability	0.061^{***}	0.008***	0.021***	0.010***	0.026***	0.079***
	(0.001)	(0.000)	(0.001)	(0.001)	(0.001)	(0.001)
Observations	481,725	481,725	481,725	481,725	155,927	481,725
Mean of dependent variable	0.033	0.005	0.009	0.008	0.020	0.043

PANEL B: Children with non-veteran parents

					Independent	
	Cognitive	Physical	Sensory	Self-Care	Living	Any
	Difficulty	Difficulty	Difficulty	Difficulty	Difficulty	Difficulty
	(1)	(2)	(3)	(4)	(5)	(6)
Parent declares a disability	0.080***	0.010***	0.032***	0.012***	0.035***	0.106^{***}
	(0.001)	(0.000)	(0.000)	(0.000)	(0.001)	(0.001)
Observations	5,126,450	$5,\!126,\!450$	$5,\!126,\!450$	$5,\!126,\!450$	1,448,513	$5,\!126,\!450$
Mean of dependent variable	0.030	0.005	0.009	0.008	0.018	0.040
p-value for H_0 : Panel A = Panel B 1	0.000	0.000	0.000	0.000	0.000	0.000

*** p < 0.01, ** p < 0.05, * p < 0.1 Standard errors clustered at the household level. Sample for column (1): all children ages 5-18; column (6) includes only individuals age 15-18. Controls: age FE, gender, dummy variables for general race categories (white, black, other - omitted; Hispanic - nonexclusive), household size FE, FE for birth order, FE for number of siblings, FE for number of grandparents in household, mother's and father's age and education FE (including indicators for missing parental information), FE for metro status, state FE, survey year FE.

disability types), despite the fact that rates of child disability among the sample of children with non-disabled parents are very similar across the two groups. This is suggestive evidence that parental disability in veteran parents is more plausibly exogenous and unrelated to underlying differences between disabled and non-disabled parents and lends credibility to the estimates discussed in Section 3.1.

That the correlations between parental disability status and child outcomes are 434 similar across the veteran and non-veteran populations speaks to the external validity of 435 our findings - namely, that the effects of parental disability are not likely limited to 436 veteran families. However, we recognize that veteran disability may differ from civilian 437 disability along key dimensions. For example, a veteran's loss of vision due to military 438 service may be very different from the lack of vision in a civilian; in particular, the 439 trauma associated with a loss of vision in wartime settings may yield additional 440 emotional and psychological consequences for veterans. This could mean that the effects 441 of veteran disability on children might be different than the effects of civilian disability. 442 Thus we regard the external validity of our results beyond the veteran population with 443 caution. 444

429

430

431

432

4 Mechanisms

In this section, we present evidence for two particular channels through which parental disability affects child outcomes: through a reduction in household economic resources and through an increase in the need to care for parents with disabilities. 446

4.1 Income

A long line of research documents the negative effects that disability has on adult labor market outcomes (see [4] for a review). Relatedly, prior work has found that parental job loss can adversely impact children (see, for example, [12] and [13]). Thus, one way in which parental disability could affect child schooling and health outcomes is through its impact on parents' ability to earn income and on household resources more generally.

To investigate this channel, we first explore the relationship between parental SCDR 455 and household income per capita in Table 7. Children with more severely disabled 456 parents live in households with significantly lower income; household income per capita 457 is 1,185 (6.7%) lower for children with severely disabled parents on net (column 1). 458 This measure of income per capita includes transfers from the Veterans Administration 459 (VA) and other assistance programs, so it suggests that overall economic resources are 460 lower for these children. In fact, when we examine other specific sources of income, we 461 find that income from "other sources" – which explicitly includes payments from the 462 VA, we see that income per capita from this source increases steeply with parental 463 disability, as expected (column 2).²⁴ Similarly, Supplemental Security Income (SSI) 464 payments also increase with parental disability, though in much smaller amounts 465 (column 3); this reflects the fact that transfers from the VA lower eligibility for other 466 types of assistance, including SSI. Welfare receipt does not seem to be systematically 467 related to parental SCDR. Column 5 illustrates that earned income (income from wages, 468 salary, and owned business and farms) is the driving force behind the lower observed 469 household incomes for children of severely disabled parents. The more severely disabled 470 the parent is, the less earned income in the household, and transfers from the VA and 471 other sources are not enough to fully offset this lost income.²⁵ 472

445

 $^{^{24}}$ Our results in Table 7 show that on average, within the group of veteran parents with an SCDR of 70 percent or higher, families experience a large decline in per capita income and a large increase in "Other Income," which includes VA payments. However, this average effect could mask substantial heterogeneity within this group, particularly for veterans with an SCDR of 100%, who receive considerably higher VA payments (see S1 Fig). For veterans with an SCDR of 100%, we might expect that the increase in "Other Income" is larger and thus the reduction in total household per capita income lower.

 $^{^{25}}$ [47] find that earnings losses for veterans with SCDRs are smaller than VA payments on average. However, we believe that our findings – specifically, that household earnings per capita are not fully offset by VA transfers – are consistent with [47] for two main reasons. First, disabled veterans often require care, which reduces the labor supply and earnings of other household members in addition to the disabled veteran herself/himself (we discuss this in more detail in Section 4.2). Thus, total household earnings per capita may fall by more than VA payments even if a veteran's own earnings losses are less than VA payments. Second, we study the sample of disabled veterans living with children. VA payments take into account household demographics but are most generous on a per capita basis for veterans without dependents (see S1 Fig). Thus even if veterans' earnings losses are less than VA payments on

	Household	Other Income	SSI	Welfare	Earned
	Income	(includes	Income	Income	Income
	per capita	VA payments)	per capita	per capita	per capita
	(1)	(2)	(3)	(4)	(5)
Parental SCDR					
10 to 20 Percent	-248.5^{***}	374.5^{***}	4.15	0.529	-907.0***
	(90.7)	(11.5)	(3.42)	(1.566)	(86.9)
30 to 40 Percent	-433.1***	1,085.0***	8.59**	0.138	-1,841.0***
	(112.1)	(17.3)	(4.35)	(2.155)	(106.6)
50 to 60 Percent	-400.1***	$2,054.5^{***}$	17.43^{***}	-2.680	$-2,978.5^{***}$
	(121.5)	(25.3)	(5.32)	(1.910)	(115.6)
70 Percent or Higher	-1,185.1***	4,458.8***	121.76^{***}	0.328	$-6,151.4^{***}$
	(85.9)	(34.7)	(6.78)	(2.021)	(82.8)
Observations	481,620	481,725	481,725	481,725	481,725
Mean of dep. var.	17796	349.1	72.81	21.02	16514
p-value for test					
that SCDR 10-20=SCDR 70+	0.000	0.000	0.000	0.934	0.000

Table 7. Degree of Parental Disability and Household Income (Veteran Sample)

*** p<0.01, ** p<0.05, * p<0.1 Standard errors clustered at the household level. Omitted group: Children in families where neither parent has a disability rating (SCDR=0). Sample: all children ages 5-18 living with a veteran parent. Controls: age FE, gender, dummy variables for single race categories (white, black, Hispanic), household size FE, FE for birth order, FE for number of siblings, FE for number of grandparents in household, mother's and father's age, education, and marital status FE as well as indicators for whether mothers and father served in 2001 and later (including indicators for missing parental information), FE for metro status, state FE, survey year FE. Mean is reported for children in families where neither parent has an SCDR. All values have been winsorized at the 99.5th percentile across the entire sample (including non-veteran households).

> In Table 8, we further document the effect of parental disability on parental labor 473 supply. Specifically, we show that, among children with a veteran father, fathers' 474 probability of work and work hours decline sharply when fathers are more severely 475 disabled (columns 1 and 2). The pattern is strikingly similar for children with a veteran 476 mother (columns 3 and 4) despite the smaller sample sizes.²⁶ The reductions in labor 477 supply are large and meaningful; severely disabled fathers are 32.8 percentage points 478 less likely to work relative to non-disabled veteran fathers (who work at high rates, 479 92.2%). The gradient is very steep; veteran fathers with a less severe disability (10-20) 480 percent) are only 1.4 percentage points less likely to work than non-disabled fathers and 481 the difference in the effects of severe disability (70 percent or more) are significantly 482 different from the effects of less severe disability (10-20 percent). 483

4.2 Caring for disabled parents

Another way in which parental disability can affect children is through the additional caregiving needs a disabled parent may require. If children devote time to caring for a disabled parent, this may decrease the time they spend on schooling activities (such as homework) and other activities (such as work for older children). This suggests that some of the adverse effects on children documented in Section 3.1 may be due to a reallocation of children's time toward parental care and away from human capital

average, this may not be true on a per capita basis for veterans with families.

 $^{^{26}}$ We display the effects of parental disability on the non-disabled parents' labor supply in S7 Table and discuss the results in Section 4.2.

	Father Works	Father's Work	Mother Works	Mother's Work
		Hours		Hours
	(1)	(2)	(3)	(4)
Parental SCDR				
10 to 20 Percent	-0.014***	-48.2***	-0.023***	-58.6***
	(0.002)	(6.9)	(0.008)	(18.6)
30 to 40 Percent	-0.034***	-134.8^{***}	-0.049***	-144.3^{***}
	(0.003)	(9.2)	(0.010)	(22.1)
50 to 60 Percent	-0.078***	-276.7^{***}	-0.097***	-267.7***
	(0.004)	(11.8)	(0.012)	(26.0)
70 Percent or Higher	-0.328***	-859.3***	-0.318***	-721.5***
	(0.004)	(10.0)	(0.011)	(21.5)
Observations	433,903	433,903	73,283	73,283
Mean of dep. var.	0.923	2030	0.775	1384
p-value for test				
that SCDR 10-20=SCDR 70+	0.000	0.000	0.000	0.000

Table 8 Degree of Parental Disability and Parental Labor Supply (Veteran Sample)

*** p<0.01, ** p<0.05, * p<0.1 Standard errors clustered at the household level. Omitted group: Children in families where neither parent has a disability rating (SCDR=0). Sample: all children ages 5-18 living with a veteran father (columns 1 and 2) or mother (columns 3 and 4). Controls: age FE, gender, dummy variables for single race categories (white, black, Hispanic), household size FE, FE for birth order, FE for number of siblings, FE for number of grandparents in household, mother's and father's age, education, and marital status FE as well as indicators for whether mothers and father served in 2001 and later (including indicators for missing parental information). FE for metro status, state FE. survey year FE. Mean is reported for children in families where neither parent has an SCDR.

> accumulation. This channel is more likely to be relevant for older children, who are more likely to be capable of providing care.

To explore this mechanism, we first study work outcomes for teens (ages 16 and 493 older, for whom the ACS contains work information). Other than schooling, work 494 outcomes are the only other type of information that the ACS collects with regard to 495 time allocation. In Table 9, we show that teens are 4.6 percentage points less likely to 496 work than teens with a non-disabled parent (column 1). This is a large effect - around 497 11% over the average work probability of teens without a disabled parent - and it is 498 statistically significant at the 1% level, as is the difference relative to teens with a less 499 disabled parent. Hours of work are also lower for teens with severely disabled parents 500 (column 2) though this appears to be driven by extensive margin changes in work status, 501 as there are no effects on hours conditional on working (column 3). Interestingly, we 502 find that working teens with more severely disabled parents have jobs that require lower 503 transit time (column 4). This is also consistent with the notion that teens that must 504 care for disabled parents have less time to for other activities, including commuting to 505 jobs. Finally, in column 5 we do not observe that parental disability is systematically 506 related to hourly earnings. We regard the results on work hours conditional on working 507 (columns 3-5) as suggestive, as we find that parental SCDR affects work status and we 508 lack a separate instrument for selection into work. 509

We find corroborating evidence of the time cost of having a disabled family member 510 in S7 Table, where we show that mothers' labor supply is adversely affected when 511 fathers are more severely disabled (and vice versa for disabled mothers). Mothers are 7.9 512 percentage points (10%) less likely to work when fathers are severely disabled, and this 513 difference is significant at the 1% level relative to families with non-disabled fathers and 514 families with less severely disabled fathers (column 1). Mother's hours of work are also 515 lower in families with a more severely disabled father (column 2). The effects are very 516

491

Table 9. Degree of Pa	arentai Disab	and reen	Labor Supply	veteran Samp	ie)
	Works	Work Hours	Work Hours	Travel Time	Hourly
		(All)	(Workers)	(minutes)	Earnings
	(1)	(2)	(3)	(4)	(5)
Parental SCDR					
10 to 20 Percent	0.003	-3.14	-13.0*	-0.230	0.165
	(0.005)	(3.99)	(7.83)	(0.260)	(0.164)
30 to 40 Percent	-0.008	4.22	17.2	0.459	-0.065
	(0.007)	(5.30)	(10.6)	(0.377)	(0.220)
50 to 60 Percent	-0.018^{**}	-4.60	8.23	-0.672^{*}	-0.503***
	(0.008)	(6.33)	(13.2)	(0.402)	(0.180)
70 Percent or Higher	-0.046***	-21.0***	-7.83	-0.836***	-0.277
	(0.006)	(4.37)	(9.85)	(0.320)	(0.175)
Observations	116,001	116,001	46,655	$46,\!655$	$46,\!652$
Mean of dep. var.	0.409	188.1	460.1	10.77	6.873
p-value for test					
that SCDR 10-20=SCDR 70+	0.000	0.001	0.662	0.122	0.049

Table 9. Degree of Parental Disability and Teen Labor Supply (Veteran Sample)

*** p<0.01, ** p<0.05, * p<0.1 Standard errors clustered at the household level. Omitted group: Children in families where neither parent has a disability rating (SCDR=0). Sample: all teenagers ages 16-18 living with a veteran parent; sample is further restricted to working children in columns 3-5. Controls: age FE, gender, dummy variables for single race categories (white, black, Hispanic), household size FE, FE for birth order, FE for number of siblings, FE for number of grandparents in household, mother's and father's age, education, and marital status FE as well as indicators for whether mothers and father served in 2001 and later (including indicators for missing parental information), FE for metro status, state FE, survey year FE. Mean is reported for children in families where neither parent has an SCDR. Earnings are winsorized at the 99.5th percentile.

similar in families with a veteran mother (columns 3-4). Overall, the evidence in Table 9 and S7 Table suggest that more severely disabled parents require additional care, and the time devoted to this care comes at the cost of both teen and spousal labor supply. 519

We further investigate the role of care for a disabled parent in Table 10. We begin 520 by reproducing the baseline relationship between parental disability status and child 521 schooling status in column 1. We see that children whose parent has a disability is 0.6 522 percentage points less likely to currently be in school. In column 2, we decompose this 523 average effect by type of parental disability. The negative relationship between parental 524 disability and schooling is strongest for types of parental disability that explicitly 525 require care – disabilities that limit an individual's ability to perform basic physical, 526 self-care, or mobility activities (i.e., physical, mobility, or self-care disabilities).²⁷ The 527 coefficient is substantially and significantly higher for parental disabilities that require 528 care (p-value for difference = 0.002). In column 3 we decompose the average effect by 529 child age. Here we see that the negative relationship is entirely driven by high 530 school-age children (ages 14-18), who are most likely to be able to provide care for 531 disabled parents; high school-age children are 1.4 percentage points less likely to 532 currently be in school if they have a disabled parent. In fact, there is no significant 533 relationship between schooling status and parental disability for elementary school-age 534 children (ages 5-10) or middle school-age children (ages 11-13). In column 4, we show 535 that on average, teens (16-18) are 3.2 percentage points less likely to work when they 536 have a disabled parent. In column 5, we show that this relationship is specific to parents 537

 $^{^{27}6.6\%}$ of children in the sample live with a parent with a disability that does not require care while 9.4% of children in the sample live with a parent with a disability that requires care.

Table 10.	Effects of Parental Disability by Type of Parental Disability and Child Age (Veteran
	Sample)

	1	,			
	In School	In School	In School	Works	Works
	(1)	(2)	(3)	(4)	(5)
Parent has any disability	-0.006***			-0.032***	
	(0.001)			(0.004)	
Parent has a disability		-0.008***			-0.050***
that requires care		(0.001)			(0.004)
Parent has a disability		-0.003***			-0.003
that does not require care		(0.001)			(0.006)
Parent has any disability		. ,	-0.000		· · · ·
\times Age 5-10			(0.001)		
Parent has any disability			0.000		
× Age 11-13			(0.001)		
Parent has any disability			-0.014***		
\times Age 14-18			(0.001)		
Observations	481,725	481,725	481,725	116,001	116,001
Mean of dep. var.	0.967	0.967	0.967	0.411	0.411
p-value for test that					
requires care $=$ does not require care		0.002	0.000^{1}		0.000

*** p<0.01, ** p<0.05, * p<0.1 Standard errors clustered at the household level. Omitted group: Children in families where neither parent reports a disability. Disabilities that require care are physical, mobility, or self-care disabilities. Sample: all children ages 5-18 living with a veteran parent; sample is further restricted to teenagers ages 16-17 in columns 4-5. Controls: age FE, gender, dummy variables for single race categories (white, black, Hispanic), household size FE, FE for birth order, FE for number of siblings, FE for number of grandparents in household, mother's and father's age, education, and marital status FE as well as indicators for whether mothers and father served in 2001 and later (including indicators for missing parental information), FE for metro status, state FE, survey year FE. Mean is reported for children in families where neither parent has an SCDR.

¹ p-value for the test that effects of parental disability for children age 5-10 is the same as for children 14-18.

who have a disability that requires care; there is no relationship between parental disability and teen labor supply when the disabled parent does not require care. 539

Taken all together, these findings are consistent with the possibility that older children spend more time caring for disabled parents and are thus unable to attend school or work at the same rates as their counterparts with non-disabled or less-disabled parents. This interpretation of the findings aligns with the limited existing evidence on the effect of parental illness on time spent in household chores and caregiving [31, 48, 49].²⁸

4.3 Locational preferences and access to health care and schooling facilities

Another possible channel through which parental disability affects human capital 548 investments in children is through its impact on the locational preferences of veterans. 549 For example, more severely disabled veterans may not be able to afford to live in 550 neighborhoods with easy access to private schools, and therefore their children may be 551 less likely to attend private schools. Alternatively, more severely disabled veterans may 552 need to locate closer to health care facilities, and thus their children may also have 553

540

541

542

543

544

545

546

 $^{^{28}}$ For example, [31] use data from China and find a strong negative association between parental chronic health conditions and disability and children's school enrollment, attendance, and performance as well as on educational spending; they find a positive (though not always statistically significant) association between maternal ill health and time spent working in the household.

easier access to health care and thus experience improved health outcomes.

To explore this possibility, we show that our results are robust to including county of 555 residence fixed effects in S7 Table. We argue that within counties, access to schooling 556 and healthcare access is more similar across households. Even once we focus on 557 within-county comparisons, we find that more children of more disabled parents are less 558 likely to attend private school and are more likely to have a disability themselves. 559

5 Conclusion

We find evidence that children face disadvantages in terms of schooling and own 561 disability outcomes when a parent is disabled. Importantly, we document that there 562 exists a gradient in child outcomes in the population of veteran families, for whom the 563 degree of parental disability is more plausibly exogenous. This appears to operate at 564 least in part through two channels. First, parental disability lowers household income 565 and thus the resources available to invest in children's human capital. Second, parents 566 with disabilities can require care, which is likely to be provided by older children in the 567 household, reducing time these teens spend on schooling and work. 568

We believe this is an important step towards a deeper understanding of the 569 relationship between parental disability and child outcomes, especially for vulnerable 570 populations. Despite the fact that more severely disabled veterans received greater 571 disability benefits, this paper shows that their children are still worse off, implying that 572 disability related social safety nets are perhaps not able to fully insure children in 573 military families. We highlight that these relationships are also likely to hold in the 574 broader population of non-veterans. However, our analysis is limited by data 575 availability; we are only able to study the effects of parental disability on a small set of 576 outcomes that are coarsely measured. Analyzing the impact of parental disability on 577 different facets of child development represents an important avenue for future research. 578

Supporting information

S1 Fig. Vetera	ns Compensation Benefits by Service Connected Disability	58
Rate and Demo	graphics.	58
S1 Table. Bala Sample).	nce on Observable Characteristics by SCDR (Veteran	58: 58:
S2 Table. Impa Disability (Vete	acts of Parental Disability for Children Born Before Parent's ran Sample).	584 585
S3 Table. Effe	cts of Any Adult SCDR (Veteran Sample).	586

554

560

S4 Table. Relationship between SCDR and number of children in the household (Veteran Sample.	587 588
S5 Table. Heterogeneity by Race and Sex of Child (Veteran Sample).	589
S6 Table. Heterogeneity by the Identity of Disabled Parent (Veteran Sample).	590 591
S7 Table. Parental Disability and the Labor Supply of the Non-Disabled Parent (Veterans Sample).	592 593
S7 Table. Effects of Parent SCDR Controlling for County Fixed Effects (Veteran Sample).	594 595
References	596
 Heckman JJ. Skill Formation and the Economics of Investing in Disadvantaged	597
Children. Science. 2006;312(5782):1900–1902.	598
 Currie J, Vogl T. Early-Life Health and Adult Circumstance in Developing	599
Countries. Annual Review of Economics. 2013;5:1–36.	600
3. Filmer D. Disability, poverty, and schooling in developing countries: results from 14 household surveys. The World Bank Economic Review. 2008;22(1):141–163.	601 602
 Haveman R, Wolfe B. The Economics of Disability and Disability Policy.	603
Handbook of Health Economics. 2000;1B:995–1051.	604
5. Hoogerveen JG. Measuring Welfare for Small but Vulnerable Groups: Poverty	605
and Disability in Uganda. Journal of African Economies. 2005;14(4):603–631.	606
 Mitra S, Posarac A, Vick BC. Disability and poverty in developing countries: a	607
snapshot from the world health survey. World Bank Social Protection Working	608
Paper. 2011;(1109).	609
 Mont D, Cuong NV. Disability and Poverty in Vietnam. The World Bank	610
Economic Review. 2011;25(2):323–359.	611
 Dimitrijević L, Jakubi BJ. The importance of early diagnosis and early physical	612
treatment of cerebral palsy. Facta Universitatis, Series: Medicine and Biology.	613
2005;12(3):119–122.	614
 Keogh BK, Becker LD. Early detection of learning problems: Questions, cautions,	615
and guidelines. Exceptional Children. 1973;40(1):5–11.	616
 Apouey B, Geoffard PY. Family income and child health in the UK. Journal of	617
health economics. 2013;32(4):715–727.	618
 Case A, Lubotsky D, Paxson C. Economic Status and Health in Childhood: The	619
Origins of the Gradient. American Economic Review. 2002;92(5):1308–1334.	620
 Stevens AH, Schaller J. Short-run effects of parental job loss on children's	621
academic achievement. Economics of Education Review. 2011;30(2):289–299.	622

13.	Lindo JM. Parental job loss and infant health. Journal of Health Economics. $2011;30(5):869-879.$	623 624
14.	Sacerdote B. How large are the effects from changes in family environment? A study of Korean American adoptees. The Quarterly Journal of Economics. 2007;122(1):119–157.	625 626 627
15.	Coneus K, Spiess CK. The intergenerational transmission of health in early childhood-Evidence from the German Socio-Economic Panel Study. Economics & Human Biology. 2012;10(1):89–97.	628 629 630
16.	Stone AL, Wilson AC. Transmission of risk from parents with chronic pain to offspring: an integrative conceptual model. Pain. 2016;157(12):2628.	631 632
17.	Hogan DP, Shandra CL, Msall ME. Family developmental risk factors among adolescents with disabilities and children of parents with disabilities. Journal of Adolescence. 2007;30(6):1001–1019.	633 634 635
18.	Charles KK. Sickness in the Family: Health Shocks and Spousal Labor Supply. Ford School of Public Policy Working Paper. 1999;00-011.	636 637
19.	Aldridge J, Becker S. Children as carers: the impact of parental illness and disability on children's caring roles. Journal of Family Therapy. 1999;21(3):303–320.	638 639 640
20.	Holmes AK, Rauch PK, Cozza SJ. When a Parent Is Injured or Killed in Combat. The Future of Children. 2013;23(2):143–162.	641 642
21.	Shandra CL, Hogan DP. The Educational Attainment Process Among Adolescents and Children of Parents with Disabilities. International Journal of Disability, Development and Education. 2009;56(4):363–379.	643 644 645
22.	Sabia JJ, Skimmyhorn WL. War! What is it Good For? The Effects of Combat Service on Economic Transitions of Veterans. Economic Self-Sufficiency Policy Research Institute Working Paper Series. 2018;(20181).	646 647 648
23.	Costa DL, Yetter N, DeSomer H. Wartime health shocks and the postwar socioeconomic status and mortality of union army veterans and their children. Journal of Health Economics. 2020;70:102281.	649 650 651
24.	Altman BM, Cooper PF, Cunningham PJ. The case of disability in the family: impact on health care utilization and expenditures for nondisabled members. The Milbank Quarterly. 1999;77(1):39–75.	652 653 654
25.	Hyatt RR, Allen SM. Disability as a "Family Affair": Parental Disability and Childhood Immunization. Medical Care. 2005;43(6):600–606.	655 656
26.	Mont D, Nguyen C. Does Parental Disability Matter to Child Education? Evidence from Vietnam. World Development. 2013;48:88–107.	657 658
27.	Currie J, Moretti E. Biology as destiny? Short-and long-run determinants of intergenerational transmission of birth weight. Journal of Labor economics. 2007;25(2):231–264.	659 660 661
28.	Elder T, Figlio D, Imberman S, Persico C. The role of neonatal health in the incidence of childhood disability. American Journal of Health Economics. 2020;6(2):216–250.	662 663 664

29.	Stabile M, Allin S. The economic costs of childhood disability. The future of children. 2012; p. 65–96.	665 666
30.	Gertler P, Levine DI, Ames M. Schooling and Parental Death. Review of Economics and Statistics. 2004;86(1):211–225.	667 668
31.	Hannum E, Sargent T, Yu S. Poverty, Parental Ill Health and Children's Access to Schooling in Rural Gansu, China. Provincial China. 2009;1(2):24.	669 670
32.	Ainsworth M, Beegle K, Koda G. The Impact of Adult Mortality and Parental Deaths on Primary Schooling in North-Western Tanzania. The Journal of Development Studies. 2005;41(3):412–439.	671 672 673
33.	Gertler P, Gruber J. Insuring consumption against illness. American economic review. 2002;92(1):51–70.	674 675
34.	Black SE, Breining S, Figlio DN, Guryan J, Karbownik K, Nielsen HS, et al. Sibling Spillovers [*] . The Economic Journal. 2020;doi:10.1093/ej/ueaa074.	676 677
35.	Ruggles S, Flood S, Foster S, Goeken R, Pacas J, Schouweiler M, et al. IPUMS USA: Version 11.0 [Dataset]. Minneapolis: University of Minnesota. 2021;.	678 679
36.	Dronkers J, Robert P. The effectiveness of public and private schools from a comparative perspective. EUI Working Paper. 2003;.	680 681
37.	Coleman JS, Hoffer T. Private and public high schools: The impact of communities. New York: Basic Books; 1987.	682 683
38.	Figlio DN, Stone JA, et al. School choice and student performance: Are private schools really better. Institute for Research on Poverty Working Paper Series. 1997;(1141-97).	684 685 686
39.	Alexander KL, Pallas AM. School sector and cognitive performance: When is a little a little? Sociology of education. 1985; p. 115–128.	687 688
40.	Willms JD. Catholic-school effects on academic achievement: New evidence from the high school and beyond follow-up study. Sociology of education. 1985; p. 98–114.	689 690 691
41.	Altonji JG, Elder TE, Taber CR. Selection on observed and unobserved variables: Assessing the effectiveness of Catholic schools. Journal of political economy. 2005;113(1):151–184.	692 693 694
42.	Evans WN, Schwab RM. Finishing high school and starting college: Do Catholic schools make a difference? The Quarterly Journal of Economics. 1995;110(4):941–974.	695 696 697
43.	Neal D. The effects of Catholic secondary schooling on educational achievement. Journal of Labor Economics. 1997;15(1, Part 1):98–123.	698 699
44.	U S Department of Education, National Center for Education Statistics. Private elementary and secondary enrollment, number of schools, and average tuition, by school level, orientation, and tuition: Selected years, 1999-2000 through 2011-12; 2021.	700 701 702 703
45.	Epple D, Figlio D, Romano R. Competition between private and public schools: testing stratification and pricing predictions. Journal of public Economics. 2004;88(7-8):1215–1245.	704 705 706

46.	Welsh-Loveman J, Perry I, Bernhardt A. Data and Methods for Estimating the Impact of Proposed Local Minimum Wage Laws. Institute for Research on Labor	707 708
	and Employment, Center on Wage and Employment Dynamics, University of	709
	California, Berkeley Working Paper Series. 2014;.	710
47.	Buddin R, Han B. Is Military Disability Compensation Adequate to Offset Civilian Earnings Losses from Service-Connected Disabilities? RAND Corporation monograph series. 2012;.	711 712 713
48.	Buck FM, Hohmann GW. Parental disability and children's adjustment. Annual	714

Review of Rehabilitation. 1983;3:203–241.
49. O'Neill AM. Normal and bright children of mentally retarded parents: The Huck Finn syndrome. Child Psychiatry and Human Development. 1985;15(4):255–268.